## Modern Regression

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Lecture 02.1.2 (v1.0.2)

# Signposting

- The previous section 02.1.1 is about interpretation of Regression in general.
- This lecture contains the mathematical content for Modern Regression - the vectorised version, which uses Matrix algebra.

### Notation, Notation, Notation

- There are several choices of convention that we have to make
- Vectors of length k are also matrices, but are they k × 1 or 1 × k?
- We use  $k \times 1$ , i.e. column vectors
- Similarly there are choices about matrix derivatives
- We use derivative with respect to a column vector as a row vector
- Some resources will have everything transposed as a consequence

### Linear algebra view of covariance

The covariance matrix of a random variable X
 Where X is a vector-valued RV with length k,
 has entries:

$$\operatorname{Cov}(\mathbf{X})_{ij} = \operatorname{Cov}(\mathbf{X}_i, \mathbf{X}_j) = \mathbb{E}[(\mathbf{X}_i - \mu_i)(\mathbf{X}_j - \mu_j)].$$

The matrix form for this is:

$$\Sigma = \mathbb{E}[(\mathbf{X} - \mathbb{E}[\mathbf{X}])(\mathbf{X} - \mathbb{E}[\mathbf{X}])^T],$$

• Where  $\mu = \mathbb{E}[X]$ .

Linear algebra view of correlation

Division by standard deviations is required to correctly generalise the scalar correlation:

$$\operatorname{Corr}(X,Y) = \frac{\mathbb{E}[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X \sigma_Y}.$$

The matrix form for correlation is:

$$\operatorname{Corr}(\mathbf{X}) = (\operatorname{diag}(\Sigma))^{-1/2} \Sigma (\operatorname{diag}(\Sigma))^{-1/2}$$

The matrix inversion is not computationally challenging because it is for a diagonal matrix. Regression is analogous to linear algebra with noise

Most problems in Linear Algebra can be seen as solving a system of linear equations:

$$XA + b = 0.$$

• Where X is an n by p matrix of data,

- ▶ A is an p by 1 matrix of coefficients,
- and  $-\mathbf{b}$  is a *n*-vector of target values.
- However, data are not usually generated from a linear model.
- We therefore typically seek the least-bad fit that we can:

$$\operatorname{argmin} ||XA + \mathbf{b}||_2^2 = \sum_{i=1}^N (\mathbf{x}_i A + b_i)^2$$

- i.e. we find A and b such that they minimise the distance (in the squared L<sub>2</sub> norm)
- Linear algebra allows this very effectively!
- Linear Algebra is therefore a very powerful way to view regression.

# Matrix form of least squares

- Consider data X' with p' features (columns) and n observations.
- Given the regression problem:

$$\mathbf{y} = \mathbf{X}' \boldsymbol{\beta}' + \mathbf{b} + \mathbf{e}$$

to find:

- $\beta'$  (a matrix dimension  $p' \times 1$ ))
- ▶ and b,
- ▶ to minimise 'error': in  $e^2 = \sum_{i=1}^n \epsilon_i^2$

## Matrix form of least squares

We construct a simpler representation by adding a constant feature:

$$\mathbf{X} = \begin{bmatrix} 1 & \mathbf{X}_{11} & \cdots & \mathbf{X}_{1p'} \\ & & \cdots \\ 1 & \mathbf{X}_{n1} & \cdots & \mathbf{X}_{np'} \end{bmatrix}$$

▶ which has p = p' + 1 features.
▶ We now solve the analogous equation:

$$\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}$$

which has the same solution but is in a more convenient form.

# Mean Squared Error (MSE)

The prediction error is:

$$\mathbf{e}(\beta) = \mathbf{y} - \mathbf{X}\beta$$

Using the notation that e is a p by 1 matrix
The estimation error is written in matrix form:

$$MSE(\beta) = \frac{1}{n} \mathbf{e}^T \mathbf{e}$$

$$\blacktriangleright$$
 Why?  $\mathbf{e}^T \mathbf{e} = \sum_{i=1}^n e_i^2$ 

 Hence MSE(β) is a 1 × 1 matrix, i.e. a scalar, and |MSE(β)| = MSE(β).

Noticing this sort of thing makes the matrix algebra easier.

We want to minimise this MSE with respect to the parameters β.

### How to do the Matrix Algebra

Lecture 13 of Cosma Shalizi's notes is a really helpful reminder!

Look at the Matrix Algebra Cheat Sheet - specifically:

- How does a transpose work?
- How do you re-order elements?
- How does a gradient work in linear and quadratic forms?

# Minimising MSE

• Taking (vector) derivatives with respect to  $\beta$ :

$$\nabla \text{MSE}(\beta) = \frac{1}{n} (\nabla \mathbf{y}^T \mathbf{y} - 2\nabla \beta^T \mathbf{X}^T \mathbf{y} + \nabla \beta^T \mathbf{X}^T \mathbf{X}\beta) \quad (1)$$
$$= \frac{1}{n} (0 - 2\mathbf{X}^T \mathbf{y} + 2\mathbf{X}^T \mathbf{X}\beta) \quad (2)$$

• which is zero at the optimum  $\hat{\beta}$ :

$$\mathbf{X}^T \mathbf{X} \hat{\boldsymbol{\beta}} - \mathbf{X}^T \mathbf{y} = 0$$

with the solution:

$$\hat{\boldsymbol{\beta}} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}.$$

Exercise: For the case p' = 1, check that this solution is the same as you can find in regular linear algebra textbooks.

### The Hat Matrix

There is an important and response independent quantity hidden in the prediction:

$$\mathbf{H} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T$$

► The fitted values are:

$$\hat{\mathbf{y}} = \mathbf{X}\hat{\boldsymbol{\beta}} = \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{y} = \mathbf{H}\mathbf{y}$$

H is dimension N × N
H "projects" y into the fitted value space ŷ

#### Properties of the Hat Matrix

- ▶ Influence:  $\frac{\partial \hat{y}_i}{\partial y_j} = H_{ij}$ . So H controls how much a change in one observation changes the estimates of each other point.
- symmetry:  $H^T = H$ . So influence is symmetric.
- Idempotency: H<sup>2</sup> = H. So the predicted value for any projected point is the predicted value itself.
- You should read up on these and other vector algebra properties.

#### Residuals and the Hat Matrix

The residuals can be written:

$$\mathbf{e} = \mathbf{y} - \mathbf{H}\mathbf{y} = (\mathbf{I} - \mathbf{H})\mathbf{y}$$

► I - H is also symmetric and idempotent, and can also be interpreted in terms of Influence.

Because of this,

$$MSE(\hat{\beta}) = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H})^T (1 - \mathbf{H}) \mathbf{y} = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H}) \mathbf{y}$$

#### Expectations

If the data were generated by our model(!) then they are described by an RV Y (an *n*-vector):

$$\mathbf{Y}_i = \mathbf{x}_i \boldsymbol{\beta} + \boldsymbol{\epsilon}_i$$

 $\blacktriangleright$   $\mathbf{x}_i$  is still a vector but *not* a Random Variable!

- $\epsilon$  is an  $n \times 1$  matrix of RVs with mean **0** and covariance  $\sigma_s^2 I$ .
- From this it is straightforward to show that the fitted values are unbiased:

$$\mathbb{E}[\hat{\mathbf{y}}] = \mathbb{E}[\mathbf{H}\mathbf{Y}] = \mathbf{x}\beta$$

 using the properties of Expectations with the symmetry and idempotency of H.

#### Covariance

#### Similarly, it is straightforward to show that

$$\operatorname{Var}[\hat{\mathbf{y}}] = \sigma_s^2 \mathbf{H}$$

using the properties of Variances with the symmetry and idempotency of  $\ensuremath{\mathrm{H}}$  .

In other words, the covariance of the fitted values is determined entirely by the structure of the covariates, via the Hat matrix.

### Implications

- Matrix form is a massive simplification of complex algebra
- It is easy to check that e.g. dimensions make sense
- These vector calculations are repeated in many machine-learning methods
- The details change but the principle remains
- Linear-Algebra loss minimisation techniques are extremely important
- They often sit inside a wider argument, e.g. updated conditional on some other parameters

## Reflection



Be able to define correlation and regression in multivariate context

Be able to perform basic calculations using these concepts

- Be able to extend intuition about their application.
- Be able to follow the reasoning in a paper where things get complicated.
- Matrix algebra is worth reading up on!
  - Describe it for example in your assessments' reflection.

# Signposting

- Make sure to look at 02.1-Regression.R
- The mathematics behind Modern Regression is analogous to the mathematics underpinning scalable Machine Learning. It is very important.
- For accessible material see Cosma Shalizi's Modern Regression Lectures (Lectures 13-14)
- Further reading in chapters 2.3 and 3.2 of The Elements of Statistical Learning: Data Mining, Inference, and Prediction (Friedman, Hastie and Tibshirani)
- Next up: 2.2 Statistical Testing