#### Modern Regression

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Lecture 02.1.2 (v1.0.2)

# **Signposting**

- $\blacktriangleright$  The previous section 02.1.1 is about interpretation of Regression in general.
- $\blacktriangleright$  This lecture contains the mathematical content for Modern Regression - the vectorised version, which uses Matrix algebra.

### Notation, Notation, Notation

- $\blacktriangleright$  There are several choices of convention that we have to make
- $\blacktriangleright$  Vectors of length *k* are also matrices, but are they  $k \times 1$  or  $1 \times k$ ?
- $\blacktriangleright$  We use  $k \times 1$ , i.e. column vectors
- $\blacktriangleright$  Similarly there are choices about matrix derivatives
- $\blacktriangleright$  We use derivative with respect to a column vector as a row vector
- $\blacktriangleright$  Some resources will have everything transposed as a consequence

## Linear algebra view of covariance

 $\blacktriangleright$  The **covariance matrix** of a random variable X  $\blacktriangleright$  Where X is a vector-valued RV with length  $k$ ,  $\blacktriangleright$  has entries:

$$
Cov(X)_{ij} = Cov(X_i, X_j) = \mathbb{E}[(X_i - \mu_i)(X_j - \mu_j)].
$$

 $\blacktriangleright$  The matrix form for this is:

$$
\Sigma = \mathbb{E}[(X - \mathbb{E}[X])(X - \mathbb{E}[X])^T],
$$

 $\blacktriangleright$  Where  $\mu = \mathbb{E}[X]$ .

Linear algebra view of correlation

 $\blacktriangleright$  Division by standard deviations is required to correctly generalise the **scalar correlation**:

$$
Corr(X,Y) = \frac{\mathbb{E}[(X - \mu_X)(Y - \mu_Y)]}{\sigma_X \sigma_Y}.
$$

**F** The **matrix form** for correlation is:

$$
Corr(X) = (\text{diag}(\Sigma))^{-1/2} \Sigma (\text{diag}(\Sigma))^{-1/2}
$$

 $\blacktriangleright$  The matrix inversion is not computationally challenging because it is for a **diagonal matrix**.

Regression is analogous to linear algebra with noise

▶ Most problems in Linear Algebra can be seen as solving a **system of linear equations:**

$$
XA + b = 0.
$$

 $\blacktriangleright$  Where  $\overline{X}$  is an *n* by *p* matrix of data,

- $\blacktriangleright$  A is an *p* by 1 matrix of coefficients,
- $\triangleright$  and  $-\overline{b}$  is a *n*-vector of target values.
- $\blacktriangleright$  However, data are not usually generated from a linear model.
- $\blacktriangleright$  We therefore typically seek the least-bad fit that we can:

$$
\text{argmin} ||XA + \mathbf{b}||_2^2 = \sum_{i=1}^N (\mathbf{x}_i A + b_i)^2
$$

- $\blacktriangleright$  i.e. we find A and b such that they minimise the distance (in the squared  $L_2$  norm)
- $\blacktriangleright$  Linear algebra allows this very effectively!
- $\blacktriangleright$  Linear Algebra is therefore a very powerful way to view regression.

# Matrix form of least squares

- $\blacktriangleright$  Consider data  $X'$  with  $p'$  features (columns) and  $n$ observations.
- $\blacktriangleright$  Given the regression problem:

$$
\mathbf{y} = X'\beta' + \mathbf{b} + \mathbf{e}
$$

 $\blacktriangleright$  to find:

- $\blacktriangleright$  *β'* (a matrix dimension  $p' \times 1$ ))  $\blacktriangleright$  and *b*,
- ightharpoonup to minimise 'error': in  $e^2 = \sum_{i=1}^n \epsilon_i^2$

## Matrix form of least squares

 $\triangleright$  We construct a simpler representation by adding a constant feature:

$$
\mathbf{X} = \begin{bmatrix} 1 & \mathbf{X}_{11} & \cdots & \mathbf{X}_{1p'} \\ & & \cdots & \\ 1 & \mathbf{X}_{n1} & \cdots & \mathbf{X}_{np'} \end{bmatrix}
$$

ightharpoonup which has  $p = p' + 1$  features.

 $\triangleright$  We now solve the analogous equation:

$$
\mathbf{y} = \mathbf{X}\boldsymbol{\beta} + \mathbf{e}
$$

 $\blacktriangleright$  which has the same solution but is in a more convenient form.

# Mean Squared Error (MSE)

**Fig.** The **prediction** error is:

$$
\mathbf{e}(\beta) = \mathbf{y} - \mathbf{X}\beta
$$

 $\blacktriangleright$  Using the notation that e is a p by 1 matrix **The estimation error is written in matrix form:** 

$$
MSE(\beta) = \frac{1}{n} \mathbf{e}^T \mathbf{e}
$$

$$
\blacktriangleright \text{ Why? } \mathbf{e}^T \mathbf{e} = \sum_{i=1}^n e_i^2
$$

Hence  $MSE(\beta)$  is a  $1 \times 1$  matrix, i.e. a scalar, and  $|\text{MSE}(\beta)| = \text{MSE}(\beta).$ 

 $\blacktriangleright$  Noticing this sort of thing makes the matrix algebra easier.

 $\triangleright$  We want to minimise this MSE with respect to the parameters *β*.

### How to do the Matrix Algebra

[Lecture 13 of Cosma Shalizi's notes](http://www.stat.cmu.edu/~cshalizi/mreg/15/lectures/13/lecture-13.pdf) is a really helpful reminder!

 $\blacktriangleright$  Look at the [Matrix Algebra Cheat Sheet](https://dsbristol.github.io/dst/coursebook/02-MatrixCheatsheet.html) - specifically:

- $\blacktriangleright$  How does a transpose work?
- ▶ How do you re-order elements?
- $\blacktriangleright$  How does a gradient work in linear and quadratic forms?

# Minimising MSE

**►** Taking (vector) derivatives with respect to *β*:

$$
\nabla \text{MSE}(\beta) = \frac{1}{n} (\nabla \mathbf{y}^T \mathbf{y} - 2 \nabla \beta^T \mathbf{X}^T \mathbf{y} + \nabla \beta^T \mathbf{X}^T \mathbf{X} \beta) \quad (1)
$$

$$
= \frac{1}{n} (0 - 2 \mathbf{X}^T \mathbf{y} + 2 \mathbf{X}^T \mathbf{X} \beta) \quad (2)
$$

**I** which is zero at the optimum  $\hat{\beta}$ :

$$
\mathbf{X}^T \mathbf{X} \hat{\beta} - \mathbf{X}^T \mathbf{y} = 0
$$

 $\blacktriangleright$  with the solution:

$$
\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}.
$$

 $\blacktriangleright$  Exercise: For the case  $p'=1$ , check that this solution is the same as you can find in regular linear algebra textbooks.

### The Hat Matrix

▶ There is an important and response independent quantity hidden in the prediction:

$$
\mathbf{H} = \mathbf{X} (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T
$$

 $\blacktriangleright$  The fitted values are:

$$
\hat{\mathbf{y}} = \mathbf{X}\hat{\beta} = \mathbf{X}(\mathbf{X}^T\mathbf{X})^{-1}\mathbf{X}^T\mathbf{y} = \mathbf{H}\mathbf{y}
$$

 $\blacktriangleright$  H is dimension  $N \times N$  $\blacktriangleright$  H "projects" **y** into the fitted value space  $\hat{y}$ 

#### Properties of the Hat Matrix

- **► Influence**:  $\frac{\partial \hat{y}_i}{\partial w_i}$  $\frac{\partial y_{i}}{\partial y_{j}}= \mathrm{H}_{ij}.$  So  $\mathrm{H}% _{ij}$  controls how much a change in one observation changes the estimates of each other point.
- $\blacktriangleright$  symmetry:  $H^T = H$ . So influence is symmetric.
- $\blacktriangleright$  **Idempotency**:  $H^2 = H$ . So the predicted value for any projected point is the predicted value itself.
- $\triangleright$  You should read up on these and other vector algebra properties.

#### Residuals and the Hat Matrix

 $\blacktriangleright$  The residuals can be written:

$$
\mathbf{e} = \mathbf{y} - H\mathbf{y} = (I - H)\mathbf{y}
$$

 $\blacktriangleright$  I – H is also symmetric and idempotent, and can also be interpreted in terms of Influence.

 $\blacktriangleright$  Because of this,

$$
\text{MSE}(\hat{\beta}) = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H})^T (1 - \mathbf{H}) \mathbf{y} = \frac{1}{n} \mathbf{y}^T (1 - \mathbf{H}) \mathbf{y}
$$

#### **Expectations**

If the data were generated by our model(!) then they are described by an RV **Y** (an *n*-vector):

$$
\mathbf{Y}_i = \mathbf{x}_i \beta + \epsilon_i
$$

 $\blacktriangleright$   $\mathbf{x}_i$  is still a vector but *not* a Random Variable!

- $\blacktriangleright$   $\epsilon$  is an  $n \times 1$  matrix of RVs with mean  $\bm{0}$  and covariance  $\sigma_s^2 \bm{I}$ .
- ▶ From this it is straightforward to show that the fitted values **are unbiased**:

$$
\mathbb{E}[\hat{\mathbf{y}}] = \mathbb{E}[H\mathbf{Y}] = \mathbf{x}\beta
$$

using the properties of Expectations with the symmetry and idempotency of H.

#### Covariance

#### $\triangleright$  Similarly, it is straightforward to show that

$$
\text{Var}[\hat{\mathbf{y}}] = \sigma_s^2 \mathbf{H}
$$

using the properties of Variances with the symmetry and idempotency of H.

 $\blacktriangleright$  In other words, the covariance of the fitted values is determined entirely by the structure of the covariates, via the Hat matrix.

## Implications

- $\blacktriangleright$  Matrix form is a massive simplification of complex algebra
- $\blacktriangleright$  It is easy to check that e.g. dimensions make sense
- $\blacktriangleright$  These vector calculations are repeated in many machine-learning methods
- $\blacktriangleright$  The details change but the principle remains
- $\blacktriangleright$  Linear-Algebra loss minimisation techniques are extremely important
- $\blacktriangleright$  They often sit inside a wider argument, e.g. updated conditional on some other parameters

## Reflection



▶ Be able to define **correlation** and regression in multivariate context

 $\blacktriangleright$  Be able to perform basic calculations using these concepts

- $\blacktriangleright$  Be able to extend intuition about their application.
- $\blacktriangleright$  Be able to follow the reasoning in a paper where things get complicated.

 $\blacktriangleright$  Matrix algebra is worth reading up on!

▶ Describe it for example in your assessments' reflection.

# **Signposting**

- ▶ Make sure to look at 02.1-Regression.R
- $\blacktriangleright$  The mathematics behind Modern Regression is analogous to the mathematics underpinning scalable Machine Learning. **It is very important**.
- ▶ For accessible material see [Cosma Shalizi's Modern Regression](http://www.stat.cmu.edu/~cshalizi/mreg/15/lectures/) [Lectures](http://www.stat.cmu.edu/~cshalizi/mreg/15/lectures/) (Lectures 13-14)
- $\blacktriangleright$  Further reading in chapters 2.3 and 3.2 of [The Elements of](https://web.stanford.edu/~hastie/Papers/ESLII.pdf) [Statistical Learning: Data Mining, Inference, and Prediction](https://web.stanford.edu/~hastie/Papers/ESLII.pdf) (Friedman, Hastie and Tibshirani)
- $\blacktriangleright$  Next up: 2.2 Statistical Testing